



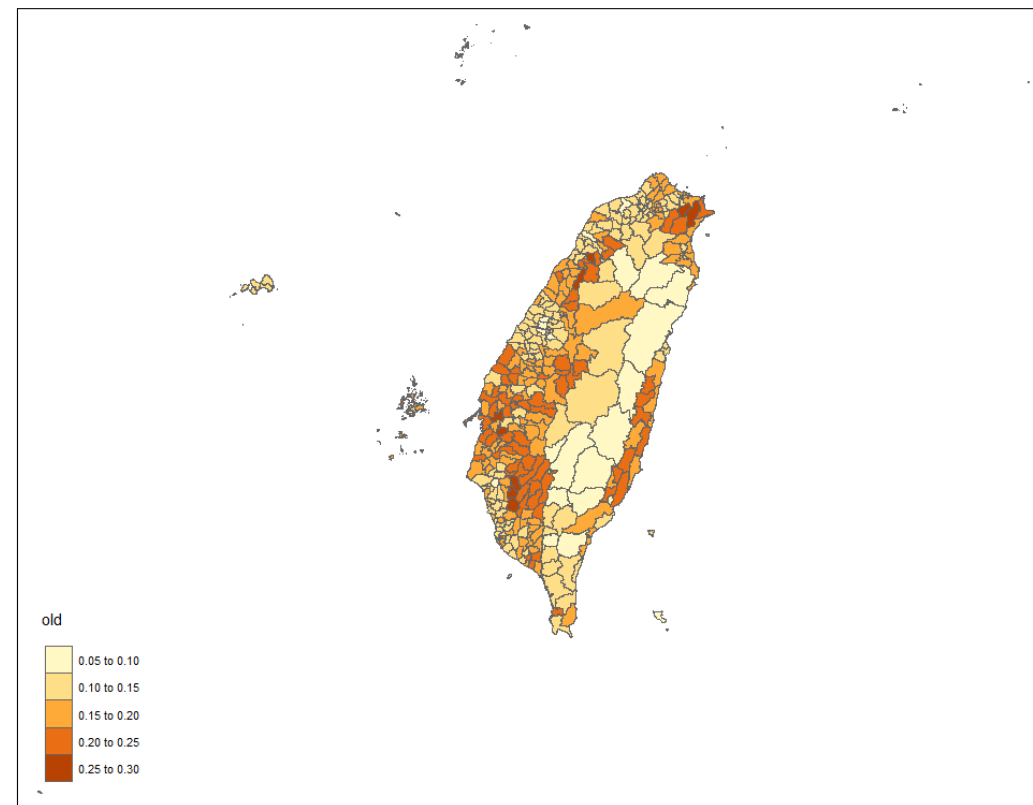
熱區分析 & 多重檢定校正

空間分析 2021.05.31
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【鄰近定義：Contiguity (Queen)】

1. 原始數值
2. LISA map
($\alpha=0.05$, 區分 HH, LL, HL, LH)
3. Standardized G_i^* values
($\alpha=0.05$, 區分 cluster, non-cluster)
4. 比較LISA進行FDR校正前後的HH熱區分布
($\alpha=0.05$, 校正前 HH)
($\alpha=0.05$, 校正後 HH)
5. 比較 G_i^* 進行Bonferroni校正前後的熱區分布
($\alpha=0.05$, 校正後 cluster)

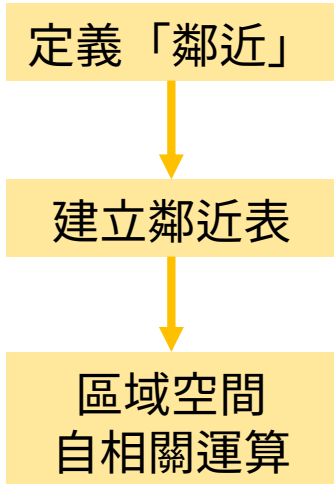
- 資料：Popn_TWN2.shp



參考答案 → → →



實作



LISA

```
TW.nb = poly2nb(TW)

TW.nb.w = nb2listw(TW.nb,
                    zero.policy=T)

LISA = localmoran(old, TW.nb.w,
                  zero.policy = T,
                  alternative = "two.sided")
```

```
> LISA
      Ii      E.Ii      Var.Ii      Z.Ii      Pr(z != 0)
220  0.8094220277 -0.025  0.17429168  1.998699187  4.564091e-02
221  0.6620073103 -0.025  0.22386090  1.452018784  1.464964e-01
222  1.3953564727 -0.025  0.17429168  3.402193655  6.684725e-04
223  0.5999538193 -0.025  0.14124553  1.662878712  9.633672e-02
224  1.5232521605 -0.025  0.14124553  4.119593286  3.795417e-05
225  1.3501517812 -0.025  0.17429168  3.293914418  9.880258e-04
226  2.3360250470 -0.025  0.14124553  6.282221450  3.337689e-10
227 -0.0299052525 -0.025  0.08616861 -0.016710399  9.866677e-01
228  0.0003684787 -0.025  0.11764114  0.073963051  9.410398e-01
229 -0.0043165576 -0.025  0.17429168  0.049543250  9.604864e-01
230 -0.0327045528 -0.025  0.06614064 -0.029958028  9.761005e-01
```

Local Moran's I
LISA[,1]

Z score
LISA[,4]

P value
LISA[,5]

Gi*

```
TW.nb = poly2nb(TW)
TW.nb.in = include.self(TW.nb)

TW.nb.w.in = nb2listw(TW.nb.in)

Gi = localG(old, TW.nb.w.in)
```

包含自己的
鄰近定義

```
> Gi
[1]  1.8911025  1.7181396  2.5357910  2.4823288
[5]  3.7590712  2.4905072  4.3849408  1.7080833
[9] -0.1426438  0.2470504  0.1209070 -1.7733190
[13] 2.4211648  2.8866465  2.4180649  2.9475747
[17] 0.9903472 -0.9465509  0.3367046 -0.9960144
[21] -1.4617826 -1.4423588 -1.6701713 -1.7999710
```

Z score of Gi*

LISA

```
LISA = localmoran(old, TW.nb.w, zero.policy = T, alternative = "two.sided" )
```

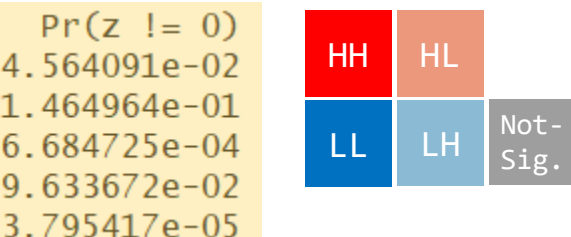
> LISA

	Ii	E.Ii	Var.Ii	Z.Ii	Pr(z > 0)
220	0.8094220277	-0.025	0.17429168	1.998699187	2.282046e-02
221	0.6620073103	-0.025	0.22386090	1.452018784	7.324819e-02
222	1.3953564727	-0.025	0.17429168	3.402193655	3.342363e-04
223	0.5999538193	-0.025	0.14124553	1.662878712	4.816836e-02
224	1.5232521605	-0.025	0.14124553	4.119593286	1.897709e-05

alternative = "greater"
預設：是否和鄰居相似(正相關)



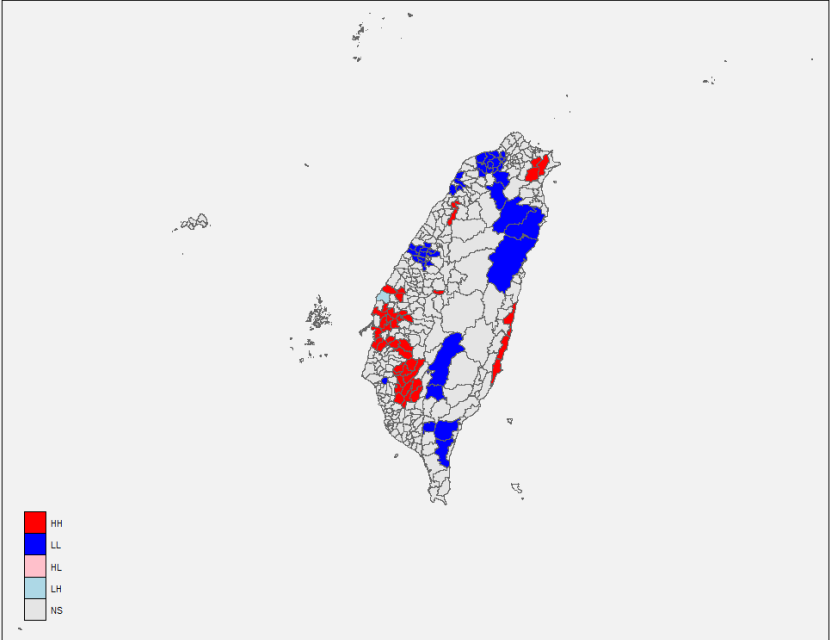
alternative = "two.sided"
我們要的：是否和鄰居有相關



```
LISA = localmoran(old, TW.nb.w, zero.policy=T, alternative ="two.sided")
```

```
z = LISA[,4]
p = LISA[,5]
diff = old - mean(old) # 自己比平均是H/L
col = c()
col[diff>0 & z>0] = "red" # H-H
col[diff<0 & z>0] = "blue" # L-L
col[diff>0 & z<0] = "pink" # H-L
col[diff<0 & z<0] = "lightblue" # L-H
col[p>0.05] = "grey90" # 不顯著
TW$colI=col
```

```
qtm(TW, 'colI')
+tm_add_legend("fill", labels=c("HH", "LL", "HL", "LH", "NS"),
col=c("red", "blue", "pink", "lightblue", "grey90"))
```



Gi*

Gi = **localG**(old, TW.nb.w.in)

※ 會列出Gi*的z分數

```

> Gi
[1] 1.8911025 1.7181396 2.5357910 2.4823288
[5] 3.7590712 2.4905072 4.3849408 1.7080833
[9] -0.1426438 0.2470504 0.1209070 -1.7733190
[13] 2.4211648 2.8866465 2.4180649 2.9475747

```

```

TW$Gi = localG(old, TW.nb.w.in)
TW$colG="grey90"
TW$colG[TW$Gi>=qnorm(.95)]="red"
qtm(TW, 'colG')

```

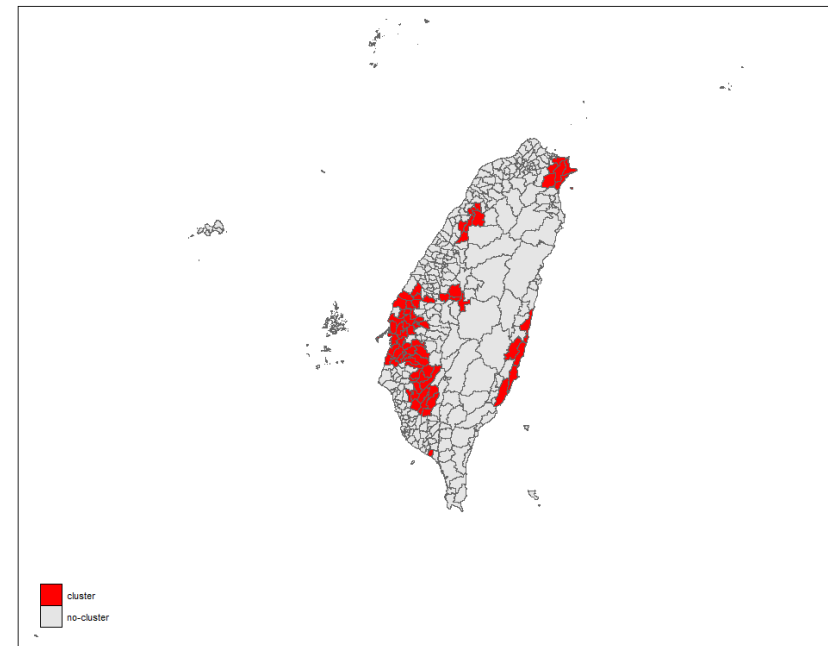
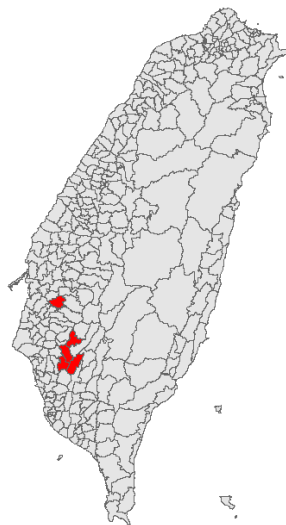
Bonferroni校正: $1-0.05/n$

`qnorm(1-0.05)` → 1.64

`qnorm(1-0.05/10)` → 3.09

`qnorm(1-0.05/100)` → 3.29

`qnorm(1-0.05/1000)` → 3.89



- Gi* 原始數值
- Gi = **localG**(old, TW.nb.w.in, **return_internals = T**)
- ※ 可以列出每個格子的Gi*, 以及期望值、變異數
- ```

> attr(Gi,"internals")

```

|   | G            | EG         | VG           |
|---|--------------|------------|--------------|
| 1 | 0.0443024793 | 0.02439024 | 1.108689e-04 |
| 2 | 0.0444890960 | 0.02439024 | 1.368440e-04 |
| 3 | 0.0510906836 | 0.02439024 | 1.108689e-04 |
| 4 | 0.0482406792 | 0.02439024 | 9.231537e-05 |

# FDR校正

```
LISA. = localmoran(old, TW.nb.w, zero.policy = T)
p = LISA.[,5]
p.adj = p.adjust(p, "fdr")

TW$colHHfdr="grey90"
TW$colHHfdr[p.adj<0.05 & diff>0]="red"
```

# FDR概念

| $i$   | $p_i$   | $p_i^*$ |
|-------|---------|---------|
| 1     | 0.00001 | 0.0010  |
| 2     | 0.00002 | 0.0010  |
| 3     | 0.00005 | 0.0017  |
| 4     | 0.0001  | 0.0025  |
| 5     | 0.0002  | 0.0040  |
| 6     | 0.0005  | 0.0083  |
| 7     | 0.001   | 0.0143  |
| 8     | 0.002   | 0.0250  |
| 9     | 0.0050  | 0.0556  |
| 10    | 0.0051  | 0.0510  |
| 11    | 0.0052  | 0.0473  |
| 12    | 0.0062  | 0.0517  |
| 13    | 0.0123  | 0.0946  |
| 14    | 0.2     | 1.4 → 1 |
| ..... | .....   | .....   |

假設共有100個樣本， $\alpha = 0.05$

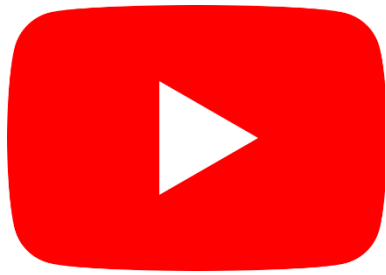
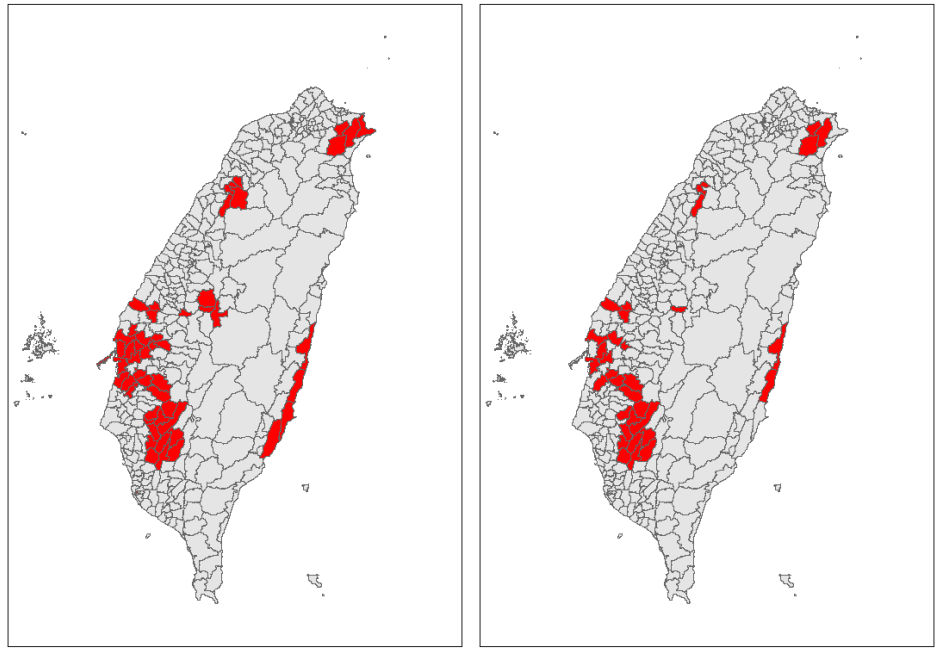
$$p_i^* = p_i \times \frac{100}{i}$$

1. 從p-value數值大的開始搜尋
2. 找到第一個熱區 (顯著)
3. 剩下的全部都是熱區

?

← 第一個熱區

Caldas de Castro, M., & Singer, B. H. (2006). Controlling the false discovery rate: a new application to account for multiple and dependent tests in local statistics of spatial association. *Geographical Analysis*, 38(2), 180-208.



多重檢定校正  
<https://youtu.be/5bqHT3Gp2W0>

## Moran's I

$$I = \frac{n \sum_i \sum_j w_{ij} (x_i - \bar{x})(x_j - \bar{x})}{W \sum_i (x_i - \bar{x})^2}$$

$$\xrightarrow{\tilde{x}_i = x_i - \bar{x}} \frac{n \sum_i \sum_j w_{ij} \tilde{x}_i \tilde{x}_j}{\sum_i \tilde{x}_i^2}$$

- $W = \sum_i \sum_j w_{ij}$
- $\sum_i (x_i - \bar{x})^2 = n \sigma_x^2 = (n-1) s_x^2 = (n-1) s_{\tilde{x}}^2$

```
> TP.nb=poly2nb(TP)
> TP.nb.w=nb2listw(TP.nb)
> M=moran.test(x,TP.nb.w)
> M$estimate[1]
Moran I statistic
-0.01261841
> TP.nb.M=nb2mat(TP.nb)
> xx=x-mean(x)
> sum(TP.nb.M*(xx%*t(xx)))/sum(xx^2)
[1] -0.01261841
> sum(TP.nb.M*(xx%*t(xx)))/(var(xx)*11)
[1] -0.01261841
```

$$I = \frac{n \sum_i \sum_j w_{ij} (x_i - \bar{x})(x_j - \bar{x})}{W \sum_i (x_i - \bar{x})^2}$$

$$= \frac{n \sum_i \sum_j w_{ij} (x_i - \bar{x})(x_j - \bar{x})}{W n \sigma^2}$$

$$= \frac{1}{W} \sum_i \sum_j w_{ij} \frac{(x_i - \bar{x})}{\sigma} \frac{(x_j - \bar{x})}{\sigma}$$

$$= \frac{1}{W} \sum_i \sum_j w_{ij} z_i z_j$$

$$= \frac{1}{W} \sum_i z_i \sum_j w_{ij} z_j = \frac{1}{W} \sum_i I_i$$

## Local Moran's I

$$I_i = z_i \sum_j w_{ij} z_j$$

$$I_i = \frac{x_i - \bar{x}}{s^2} \sum_{j \neq i} w_{ij} (x_j - \bar{x}) = z_i \sum_j w_{ij} z_j$$

- $z_i = \frac{x_i - \bar{x}}{\sigma}$
- $z_i = \frac{x_i - \bar{x}}{s}$

```
> LISA=localmoran(x,TP.nb.w)
> LISA[1]; sum(LISA[,1])/12
[1] 0.005094452 [1] -0.01261841
> z=(x-mean(x))/(sd(x)*sqrt(11/12))
> z[1]*sum(TP.nb.M[1,]*z)
[1] 0.005094452
> LISA=localmoran(x,TP.nb.w,mlvar=F)
> LISA[1]
[1] 0.004669914
> z=(x-mean(x))/sd(x)
> z[1]*sum(TP.nb.M[1,]*z)
[1] 0.004669914
```

補充：用矩陣方法一次求得所有  $I_i$

> z\*(TP.nb.M%\*z)

P.S.

$$I_i = \frac{x_i - \bar{x}}{s_i^2} \sum_{j \neq i} w_{ij} (x_j - \bar{x}); s_i^2 = \frac{\sum_{j \neq i} w_{ij} (x_j - \bar{x})^2}{n-1}$$

```
> lx=xx[1]*sum(TP.nb.M[1,]*xx)
> si2=var(x[-1])*10/11
> lx/si2
[1] 0.004670523
矩陣方法：
> xx*(TP.nb.M%*xx)/sapply(1:12,
function(i) var(x[-i])*10/11)
```

## Getis-Ord General G

$$G = \frac{\sum_i \sum_j w_{ij} x_i x_j}{\sum_i \sum_j x_i x_j}, j \neq i$$

$$\left( \text{當 } w_{ii} = 0 \xrightarrow{\text{ignore } j=i} G = \frac{\sum_i \sum_j w_{ij} x_i x_j}{\sum_i \sum_j x_i x_j - \sum_i x_i^2} \right)$$

```
> G=globalG.test(x,TP.nb.w)
> G$estimate[1]
Global G statistic
0.09243927
> G.num=sum(TP.nb.M*(x%*t(x)))
> G.den=sum(x%*t(x))-sum(x^2)
> G.num/G.den
[1] 0.09243927
```

## R package - spdep

$$I = \frac{n}{\sum_{i=1}^n \sum_{j=1}^n w_{ij}} \frac{\sum_{i=1}^n \sum_{j=1}^n w_{ij} (x_i - \bar{x})(x_j - \bar{x})}{\sum_{i=1}^n (x_i - \bar{x})^2}$$

$$I_i = \frac{(x_i - \bar{x})}{\sum_{k=1}^n (x_k - \bar{x})^2 / (n-1)} \sum_{j=1}^n w_{ij} (x_j - \bar{x})$$

## localmoran(mlvar=TRUE)

mlvar: values of local Moran's I are reported using the variance of the variable of interest (sum of squared deviances over n), but can be reported as the sample variance, dividing by (n-1) instead

## Getis-Ord Gi\*

$$G_i^* = \frac{\sum_j w_{ij} x_j}{\sum_j x_j}$$

## Getis-Ord Gi

$$G_i = \frac{\sum_j w_{ij} x_j}{\sum_j x_j}, j \neq i$$

```
> Gi.=localG(x,TP.nb.w,in,return_internals=T)
> attr(Gi.,"internals")[,1]
0.0862 0.0885 0.0923 0.0868 0.0845
> TP.nb.M.in%*x/sum(x)
0.0862 0.0885 0.0923 0.0868 0.0845
> Gi=localG(x,TP.nb.w,return_internals=T)
> attr(Gi,"internals")[,1]
0.0946 0.0966 0.0969 0.0948 0.0948
> TP.nb.M%*x/(sum(x)-x)
0.0946 0.0966 0.0969 0.0948 0.0948
```

>> 套件函數

>> 手動計算



空間自相關計算

<https://youtu.be/gOuFIxk8oFI>

14:38 ~ 46:20